



Chapter 7

Randomization

Algorithm Theory
WS 2018/19

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Randomization

Randomized Algorithm:

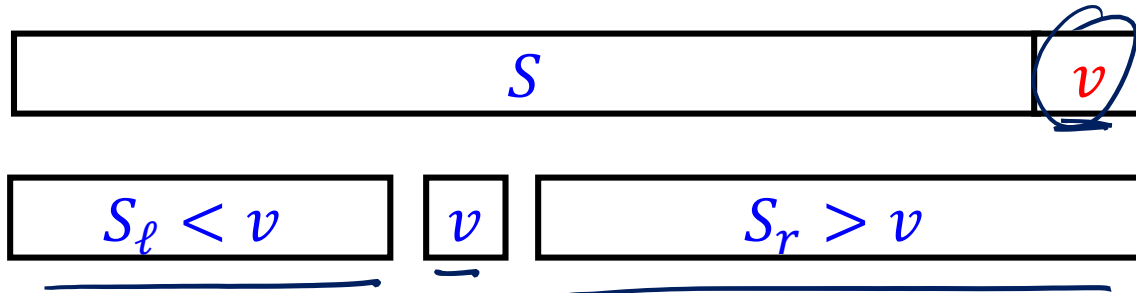
- An algorithm that uses (or can use) **random coin flips** in order to make decisions

We will see: **randomization** can be a **powerful tool** to

- Make algorithms **faster**
- Make algorithms **simpler**
- Make the analysis simpler
 - Sometimes it's also the opposite...
- Allow to **solve problems (efficiently)** that cannot be solved (efficiently) without randomization
 - True in some computational models (e.g., for distributed algorithms)
 - Not clear in the standard sequential model

Randomized Quicksort

Quicksort:



function Quick (S : sequence): sequence;

{returns the sorted sequence S }

begin

if $\#S \leq 1$ **then return** S

else { choose pivot element v in S ;

partition S into S_ℓ with elements $< v$,

and S_r with elements $> v$

return Quick(S_ℓ) v Quick(S_r)

end;

Randomized Quicksort Analysis

Randomized Quicksort: pick uniform random element as **pivot**

Running Time of sorting n elements:

- Let's just count the number of comparisons
- In the partitioning step, all $n - 1$ non-pivot elements have to be compared to the pivot

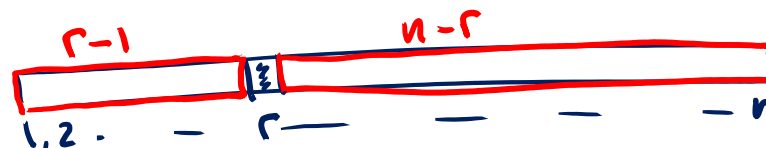
- **Number of comparisons:**

*depends on partition
random variable*

$n - 1$ + #comparisons in recursive calls

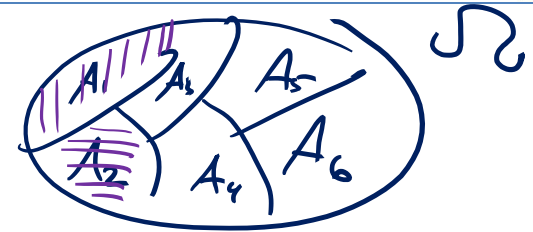
- If rank of pivot is r :

recursive calls with $r - 1$ and $n - r$ elements



Law of Total Expectation

- Given a **random variable** X and
- a set of events A_1, \dots, A_k that **partition** Ω
 - E.g., for a second **random variable** Y , we could have



$$A_i := \{\omega \in \Omega : Y(\omega) = i\}$$

Law of Total Expectation

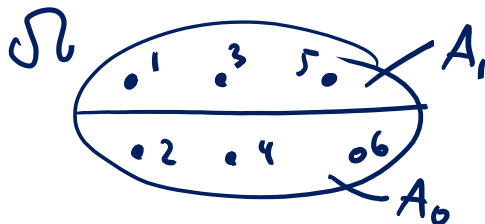
$$\underline{\mathbb{E}[X]} = \sum_{i=1}^k \underline{\mathbb{P}(A_i) \cdot \mathbb{E}[X | A_i]} = \sum_y \underline{\mathbb{P}(Y = y) \cdot \mathbb{E}[X | Y = y]}$$

\mathbb{R} (with an arrow pointing to the sum over y)

Example:

- X : outcome of rolling a die
- $A_0 = \{X \text{ is even}\}$, $A_1 = \{X \text{ is odd}\}$

$$\mathbb{E}[X] = 3,5$$



$$\mathbb{E}[X] = \mathbb{P}(A_0) \cdot \mathbb{E}[X | A_0] + \mathbb{P}(A_1) \cdot \mathbb{E}[X | A_1]$$

$$\frac{1}{2} \cdot 4 + \frac{1}{2} \cdot 3 = 3,5$$

Randomized Quicksort Analysis

Random variables: $E[C] = E[n - 1 + C_\ell + C_r]$

- C : total number of comparisons (for a given array of length n)
- R : rank of first pivot $R \in \{1, \dots, n\}$
- C_ℓ, C_r : number of comparisons for the 2 recursive calls

$$\underline{E[C]} = \underline{n - 1} + \underline{E[C_\ell]} + \underline{E[C_r]}$$

Law of Total Expectation:

$$\begin{aligned} \underline{E[C]} &= \sum_{r=1}^n \mathbb{P}(R = r) \cdot \underline{E[C|R = r]} \\ &= \sum_{r=1}^n \mathbb{P}(R = r) \cdot (n - 1 + \underbrace{E[C_\ell|R = r]}_{\text{Sorting an array of length } r-1} + \underbrace{E[C_r|R = r]}_{\text{Sorting an array of length } n-r}) \end{aligned}$$

Randomized Quicksort Analysis

We have seen that:

$$\mathbb{P}(R=r) = \frac{1}{n}$$

$$\underbrace{\mathbb{E}[C]}_{T(n)} = \sum_{r=1}^n \mathbb{P}(R=r) \cdot (n-1 + \underbrace{\mathbb{E}[C_\ell | R=r]}_{T(r-1)} + \underbrace{\mathbb{E}[C_r | R=r]}_{T(n-r)})$$

Define:

- **$T(n)$** : expected number of comparisons when sorting n elements

$$\mathbb{E}[C] = T(n)$$

$$\mathbb{E}[C_\ell | R=r] = T(r-1)$$

$$\mathbb{E}[C_r | R=r] = T(n-r)$$

Recursion:

$$\underline{T(n)} = \sum_{r=1}^n \frac{1}{n} \cdot (n-1 + T(r-1) + T(n-r))$$

$$T(0) = T(1) = 0$$

Randomized Quicksort Analysis

Theorem: The expected number of comparisons when sorting n elements using randomized quicksort is $T(n) \leq \underline{2n \ln n}$.

Proof: (by induction on n)

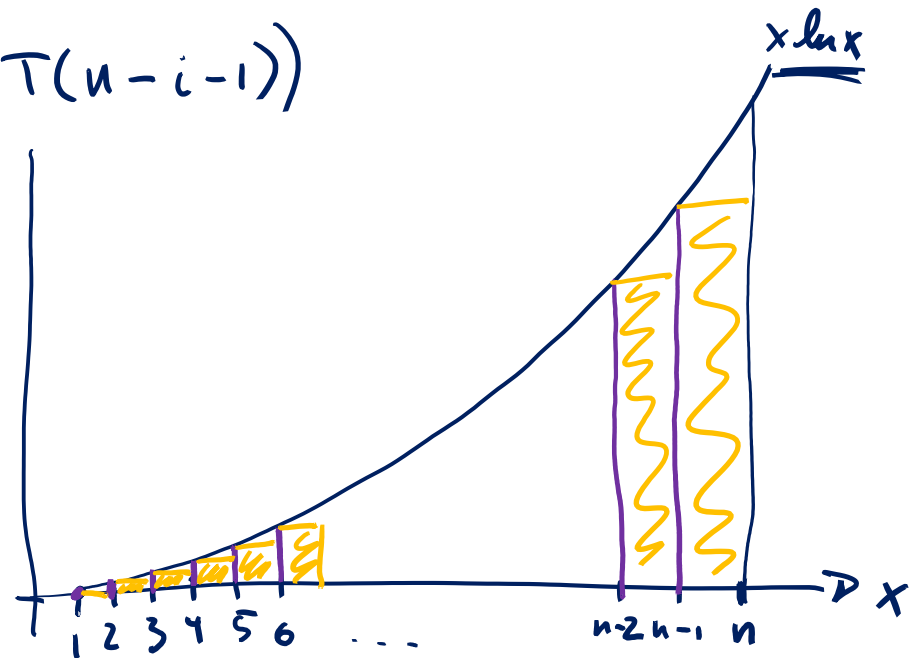
$$T(n) = \sum_{r=1}^n \frac{1}{n} \cdot (n - 1 + T(r - 1) + T(n - r)), \quad \begin{array}{l} T(1) = 0 \\ T(0) = 0 \end{array}$$

$$= n - 1 + \frac{1}{n} \cdot \sum_{i=0}^{n-1} (T(i) + T(n - i - 1))$$

$$= n - 1 + \frac{2}{n} \cdot \sum_{i=1}^{n-1} T(i)$$

$$\stackrel{\text{(I.H.)}}{\leq} n - 1 + \frac{4}{n} \cdot \sum_{i=1}^{n-1} \underline{i \cdot \ln(i)}$$

$$\leq n - 1 + \frac{4}{n} \int_1^n x \ln(x) dx$$



Randomized Quicksort Analysis

Theorem: The expected number of comparisons when sorting n elements using randomized quicksort is $T(n) \leq 2n \ln n$.

Proof:

$$T(n) \leq n - 1 + \frac{4}{n} \cdot \int_1^n x \ln x \, dx$$

$$\underline{\underline{T(n)}} \leq n - 1 + \frac{4}{n} \left[\frac{n^2 \ln n}{2} - \frac{n^2}{4} + \frac{1}{4} \right]$$

$$= n - 1 + 2n \ln n - n + \frac{1}{n}$$

$$= 2n \ln n + \underbrace{\left(\frac{1}{n} - 1 \right)}_{< 0} \leq \underline{\underline{2n \ln n}} \quad \square$$

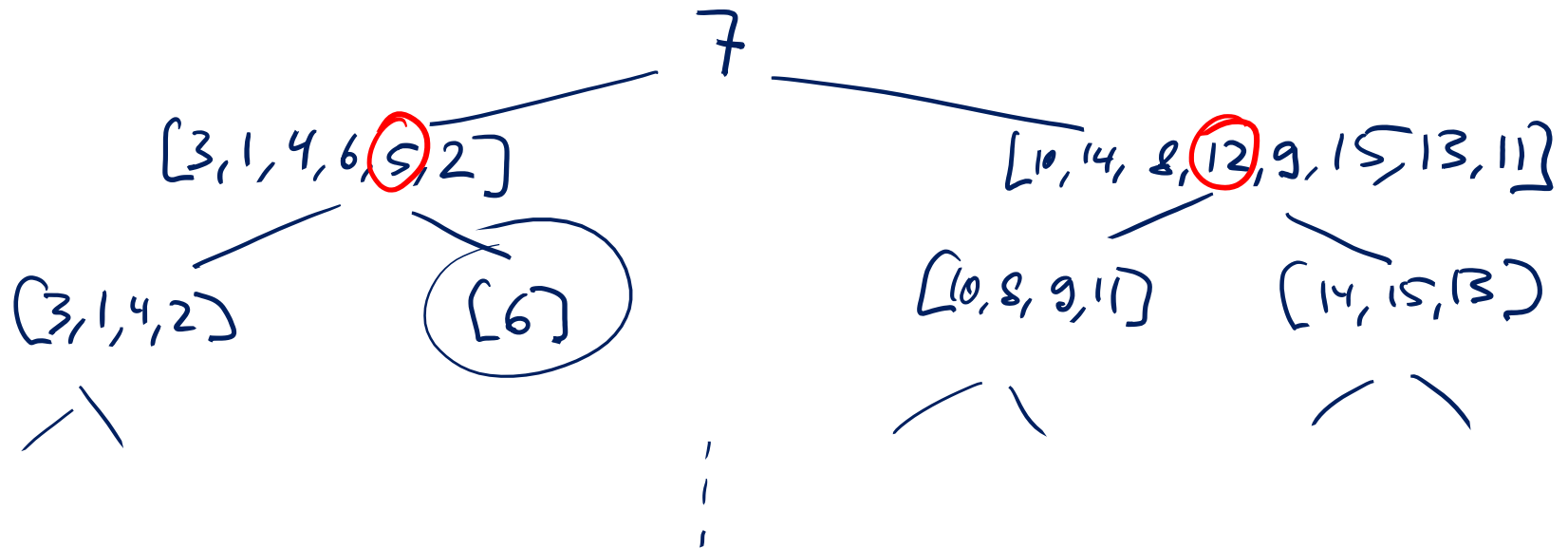
$$\Rightarrow \underline{\underline{E[C] \leq 2n \ln n}}$$

$$\int x \ln x \, dx = \frac{x^2 \ln x}{2} - \frac{x^2}{4}$$

Alternative Analysis

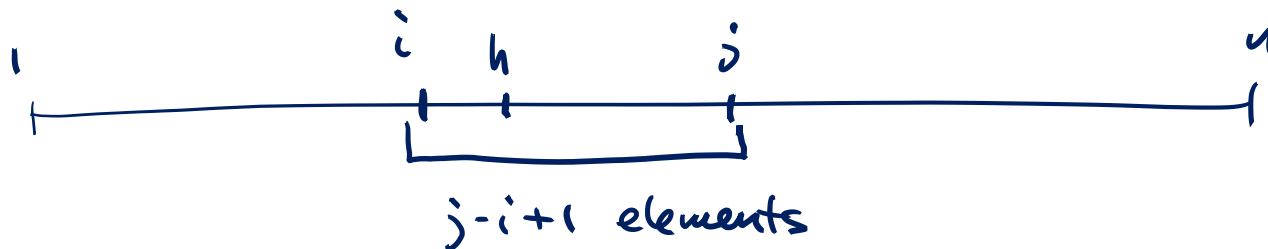
Array to sort: [7, 3, 1, 10, 14, 8, 12, 9, 4, 6, 5, 15, 2, 13, 11]

Viewing quicksort run as a **tree**:



Comparisons

- Comparisons are only between pivot and non-pivot elements
- Every element can only be the pivot once:
 - every 2 elements can only be compared once!
- W.l.o.g., assume that the elements to sort are $1, 2, \dots, n$
- Elements i and j are compared if and only if either i or j is a pivot before any element $h: i < \underline{h} < j$ is chosen as pivot
 - i.e., iff i is an ancestor of j or j is an ancestor of i



$$\mathbb{P}(\text{comparison betw. } \underline{i} \text{ and } \underline{j}) = \frac{2}{\underline{j - i + 1}}$$

Counting Comparisons

Random variable for every pair of elements (i, j) :

$$\underline{X_{ij}} = \begin{cases} 1, & \text{if there is a comparison between } i \text{ and } j \\ 0, & \text{otherwise} \end{cases}$$

$$\mathbb{P}(X_{ij}=1) = \frac{2}{j-i+1} \quad \mathbb{E}[X_{ij}] = \frac{2}{j-i+1}$$

Number of comparisons: X

$$\underline{X = \sum_{i < j} X_{ij}}$$

- What is $\mathbb{E}[X]$?

Randomized Quicksort Analysis

Theorem: The expected number of comparisons when sorting n elements using randomized quicksort is $T(n) \leq \underline{2n \ln n}$.

Proof:

- **Linearity of expectation:**

For all random variables X_1, \dots, X_n and all $a_1, \dots, a_n \in \mathbb{R}$,

$$\mathbb{E} \left[\sum_i^n a_i X_i \right] = \sum_i^n a_i \mathbb{E}[X_i].$$

$$\begin{aligned}
 X &= \sum_{i < j} X_{ij} & \mathbb{E}[X] &= \mathbb{E} \left[\sum_{i < j} X_{ij} \right] \\
 & & &= \sum_{i < j} \mathbb{E}[X_{ij}] \\
 & & &= \sum_{i < j} \frac{2}{j-i+1} = \sum_{i=1}^{n-1} \sum_{j=i+1}^n \frac{2}{j-i+1}
 \end{aligned}$$

Randomized Quicksort Analysis

Theorem: The expected number of comparisons when sorting n elements using randomized quicksort is $T(n) \leq 2n \ln n$.

Proof:

$$\begin{aligned} \mathbb{E}[X] &= 2 \sum_{i=1}^{n-1} \sum_{j=i+1}^n \frac{1}{j-i+1} = 2 \sum_{i=1}^{n-1} \sum_{k=2}^{n-i+1} \frac{1}{k} \\ &\leq 2 \sum_{i=1}^{n-1} \underbrace{\sum_{k=2}^n \frac{1}{k}}_{H(n)-1} \\ &= 2(n-1) \underbrace{(H(n)-1)}_{\leq \ln n} \\ &< \underline{2n \ln n} \quad \square \end{aligned}$$

Harmonic series:

$$H(n) := \sum_{i=1}^n \frac{1}{i}$$

$$H(n) \leq 1 + \ln n$$

Quicksort: High Probability Bound

- We have seen that the number of comparisons of randomized quicksort is $O(n \log n)$ in expectation.
- Can we also show that the number of comparisons is $O(n \log n)$ with high probability? $1 - \frac{1}{n^c}$

- **Recall:**

On each recursion level, each pivot is compared once with each other element that is still in the same “part”

Counting Number of Comparisons

- We looked at 2 ways to count the number of comparisons
 - recursive characterization of the expected number
 - number of different pairs of values that are compared

Let's consider yet another way:

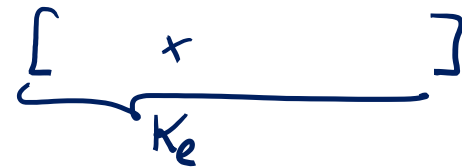
- Each comparison is between a pivot and a non-pivot
- How many times is a specific array element x compared as a non-pivot?

Value x is compared as a non-pivot to a pivot once in every recursion level until one of the following two conditions apply:

1. x is chosen as a pivot
2. x is alone

#comp. of x as non-pivot
 = depth when x becomes
 pivot/alone

Successful Recursion Level $\underline{K_1 = n}$

- Consider a specific recursion level ℓ 
- Assume that at the beginning of recursion level ℓ , element x is in a sub-array of length K_ℓ that still needs to be sorted.
- If x has been chosen as a pivot before level ℓ , we set $K_\ell := 1$

Definition: We say that recursion level ℓ is successful for element x iff the following is true:

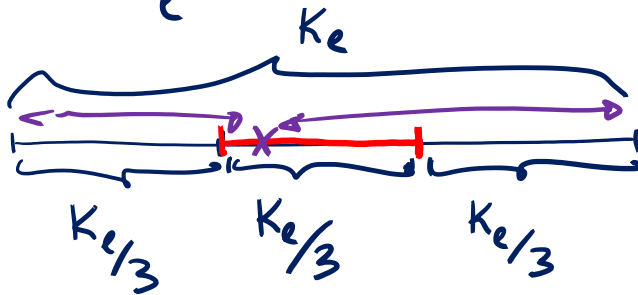
$$\underline{\underline{K_{\ell+1} = 1}} \quad \text{or} \quad \underline{K_{\ell+1} \leq \frac{2}{3} \cdot K_\ell}$$



Successful Recursion Level

Lemma: For every recursion level ℓ and every array element x , it holds that level ℓ is successful for x with probability at least $\frac{1}{3}$, independently of what happens in other recursion levels.

Proof: $k_\ell > 1$



if pivot is in the middle part

\Rightarrow both remaining arrays have
size $\leq \frac{2}{3}k_\ell \Rightarrow k_{\ell+1} \leq \frac{2}{3} \cdot k_\ell$

\Rightarrow probability for this $\geq \frac{1}{3}$

$k_\ell = 1$

$\hookrightarrow k_{\ell+1} = 1 \quad \checkmark$

Number of Successful Recursion Levels

Lemma: If among the first ℓ recursion levels, at least $\log_{3/2}(n)$ are successful for element x , we have $K_{\ell+1} = 1$.

Proof:

$$K_i = n, \quad K_{i+1} \leq K_i \quad \text{if level } i \text{ succ.} : K_{i+1} \leq \frac{2}{3} \cdot K_i$$

$$K_{\ell+1} \leq n \left(\frac{2}{3}\right)^{\# \text{ succ. levels}} \leq n \left(\frac{2}{3}\right)^{\log_{3/2} n} = n \cdot \frac{1}{n} = \underline{\underline{1}}$$

Chernoff Bounds

- Let X_1, \dots, X_n be independent 0-1 random variables and define $p_i := \mathbb{P}(X_i = 1)$.
 if $p_i = p$ for all i $X \sim \text{Bin}(n, p)$
- Consider the random variable $X = \sum_{i=1}^n X_i$
- We have $\mu := \mathbb{E}[X] = \sum_{i=1}^n \mathbb{E}[X_i] = \sum_{i=1}^n p_i$

Chernoff Bound (Lower Tail):
($\delta < 1$)

$$\mathbb{P}(X < \mu/2) < e^{-\mu/3}$$

$$\forall \delta > 0: \mathbb{P}(X < (1 - \delta)\mu) < e^{-\delta^2 \mu / 2}$$

Chernoff Bound (Upper Tail):

$$\forall \delta > 0: \mathbb{P}(X > (1 + \delta)\mu) < \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^\mu < e^{-\delta^2 \mu / 3}$$

holds for $\delta \leq 1$

Chernoff Bounds, Example

Assume that a fair coin is flipped n times. What is the probability to have

$$p_i = \frac{1}{2} \quad \mu = E[X] = n/2$$

1. less than $n/3$ heads?

$$P(X < n/3) = P(X < (1 - \frac{1}{3}) \cdot \frac{n}{2}) < e^{-\frac{1}{9} \cdot \frac{1}{2} \cdot \frac{n}{2}} = e^{-n/36}$$

2. more than $0.51n$ tails?

$$P(X > (1 + 0.02) \cdot \frac{n}{2}) < e^{-\frac{0.02^2}{3} \cdot \frac{n}{2}}$$

$$\boxed{e^{-s^2 \mu / 2}}$$

3. less than $n/2 - \sqrt{c \cdot n \ln n}$ tails?

$$P(X < (1 - \frac{\sqrt{c \cdot n \ln n}}{n}) \cdot \frac{n}{2}) \leq e^{-\frac{4c \cdot n \cdot \ln n}{n^2 \cdot 2} \cdot \frac{n}{2}} = e^{-c \ln n} = \frac{1}{n^c}$$

$$\text{w.h.p. } \# \text{ tails} / \# \text{ heads} = \frac{n}{2} \pm \underline{\underline{O(\sqrt{n \log n})}}$$

Proof of Chernoff Bound

- Independent Bernoulli random variables X_1, X_2, \dots, X_n
- $\mathbb{P}(X_i = 1) \geq p_i, X := \sum_{i=1}^n X_i, \mu := \sum_{i=1}^n p_i \geq \mathbb{E}[X]$

Chernoff Lower Tail: $\mathbb{P}(X < (1 - \delta)\mu) < e^{-\delta^2\mu/2}$

Recall

- Markov Inequality: Given non-negative rand. var. $Z \geq 0$

$$\forall t > 0: \mathbb{P}(Z > t) < \frac{\mathbb{E}[Z]}{t} \quad \mathbb{P}(Z > c \cdot \mathbb{E}[Z]) < \frac{1}{c}$$

- Independent random variables Y, Z :

$$\underline{\mathbb{E}[Y \cdot Z] = \mathbb{E}[Y] \cdot \mathbb{E}[Z]}$$

Proof of Chernoff Bound $s > 0$

$$a > b \\ e^{sa} > e^{sb}$$



- $\mathbb{P}(X_i = 1) \geq p_i$, $X := \sum_{i=1}^n X_i$, $\mu := \sum_{i=1}^n p_i \geq \mathbb{E}[X]$

$$X = \sum_{i=1}^n X_i$$

Chernoff Lower Tail: $\mathbb{P}(X < (1 - \delta)\mu) < e^{-\delta^2 \mu / 2}$

$$\mathbb{P}(X < (1 - \delta)\mu) = \mathbb{P}(-X > -(1 - \delta)\mu)$$

$$\mathbb{P}(Z > t) < \frac{\mathbb{E}[Z]}{t}$$

$$= \mathbb{P}(\underbrace{e^{-sX}}_Z > e^{-s(1 - \delta)\mu})$$

$$\stackrel{\text{(Markov)}}{<} \frac{\mathbb{E}[e^{-sX}]}{e^{-s(1 - \delta)\mu}}$$

$$\mathbb{E}[e^{-sX}] = \mathbb{E}[e^{-s \sum X_i}] = \mathbb{E}\left[\prod_{i=1}^n e^{-sX_i}\right] \stackrel{\text{indep. of } X_i}{=} \prod_{i=1}^n \mathbb{E}[e^{-sX_i}]$$

$$\mathbb{E}[e^{-sX_i}] = \mathbb{P}(X_i = 1) \cdot e^{-s \cdot 1} + \mathbb{P}(X_i = 0) \cdot e^{-s \cdot 0} = \mathbb{P}(X_i = 1) (\underbrace{e^{-s} - 1}_{< 0}) + 1 \\ \leq 1 + p_i (e^{-s} - 1)$$

Proof of Chernoff Bound $s > 0$

- $\mathbb{P}(X_i = 1) \geq p_i, X := \sum_{i=1}^n X_i, \mu := \sum_{i=1}^n p_i \geq \mathbb{E}[X]$

Chernoff Lower Tail: $\mathbb{P}(X < (1 - \delta)\mu) < e^{-\delta^2 \mu / 2}$

$$\mathbb{P}(X < (1 - s)\mu) < \frac{\mathbb{E}[e^{-sX}]}{e^{-s(1-s)\mu}} \leq e^{\mu(e^{-s} - 1 + s(1-s))}$$

$$\mathbb{E}[e^{-sX}] = \prod_{i=1}^n \mathbb{E}[e^{-sX_i}] \leq \prod_{i=1}^n (1 + p_i(e^{-s} - 1))$$

$$\begin{aligned} 1 + x &\leq e^x \\ &\leq \prod_{i=1}^n e^{p_i(e^{-s} - 1)} \\ &= e^{\sum_{i=1}^n p_i(e^{-s} - 1)} = e^{\mu(e^{-s} - 1)} \end{aligned}$$

set $s = \delta$

$$\mathbb{P}(X < (1 - \delta)\mu) < e^{\mu(e^{-\delta} - 1 + \delta(1 - \delta))} = e^{\mu(e^{-\delta} - 1 + \delta - \delta^2)}$$

Proof of Chernoff Bound

- $\mathbb{P}(X_i = 1) \geq p_i, X := \sum_{i=1}^n X_i, \mu := \sum_{i=1}^n p_i \geq \mathbb{E}[X]$

Chernoff Lower Tail: $\mathbb{P}(X < (1 - \delta)\mu) < \underline{e^{-\delta^2\mu/2}}$

$$\mathbb{P}(X < (1 - \delta)\mu) < e^{\mu(e^{-\delta} - 1 + \delta - \delta^2)}$$

$$= e^{\mu(\cancel{1 - \delta} + \frac{\delta^2}{2} - \cancel{1 + \delta} - \delta^2)}$$

$$= e^{\mu(-\frac{\delta^2}{2})}$$

_____ \square

$$e^{-\delta} = 1 - \delta + \frac{\delta^2}{2} \underbrace{[\dots]}_{< 0}$$

($\delta \leq 1$)
 $\leq 1 - \delta + \frac{\delta^2}{2}$

Number of Comparisons for x

Lemma: For every array element x , with high probability, as a non-pivot, x is compared to a pivot at most $O(\log n)$ times.

Proof:

$$\downarrow \quad \left(\frac{k}{3} \right) \quad \frac{k}{6}$$

$$\frac{e^{-k/8}}{1} = \frac{1}{n^c}$$

$$k = 8c \ln n \quad \frac{1}{n^{c-1}}$$